

# CS 49 Lecture Notes

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## 1 Counting Probabilistically

If we wanted to count the number of items in a stream (i.e.  $F_0$ ) of  $n$  items, we would need  $c = \log_2 n$  bits to store it exactly. If we allow approximation to the count, we can count  $c$  instead to keep track of the number of bits that make up  $n$ . Then, we would need a total of  $O(\log c) = O(\log \log n)$  bits. However, we cannot track  $c$  deterministically! Thus, we need a way to approximate  $c$  without using too much space.

### 1.1 Morris Counter

Morris Counter is one such way of counting  $c$  probabilistically. The algorithm is as follows:

1. Set  $c = 0$
2. When an item arrives, increment  $c$  with probability  $2^{-c}$
3. Return  $z = 2^c - 1$

Note that the mechanism to increment  $c$  can be implemented with  $O(\log c)$  space plus a random bit (used as a coin), since we can toss the coin  $c$  times and increment  $c$  only when all of them land head.

#### 1.1.1 Analysis

Let  $z(n)$  be the state of the counter after  $n$  items arrived. Deterministically,

$$z(0) = 0 \text{ and } z(1) = 1 \tag{1}$$

Since we're incrementing  $c$  with a coin toss,

$$[z(n+1)|c(n) = j] = \begin{cases} 2^j & \text{with probability } 1 - \frac{1}{2^j} \\ 2^{j+1} & \text{with probability } \frac{1}{2^j} \end{cases}$$

Therefore,  $\mathbb{E}[z(n+1)|c(n) = j] = 2^j \cdot (1 - 1/2^j) + 2^{j+1} \cdot (1/2^j) = 2^j + 1$  and

$$\begin{aligned} \mathbb{E}[z(n+1)] &= \sum_j P[c(n) = j] \mathbb{E}[z(n+1)|c(n) = j] \\ &= P[c(n) = j](2^j + 1) \\ &= 1 + \mathbb{E}[2^{c(n)}] \\ &= 1 + \mathbb{E}[z(n)] \end{aligned} \tag{2}$$

Then, iterating over the values of  $n$ ,

$$\begin{aligned}
\mathbb{E}[z(n)] &= 1 + \mathbb{E}[z(n-1)] \\
&= 1 + (1 + \mathbb{E}[z(n-2)]) \\
&\dots \\
&= n - 1 + \mathbb{E}[z(1)] \\
&= n \quad (\because \mathbb{E}[z(1)] = 1)
\end{aligned} \tag{3}$$

Therefore, the counter is an unbiased estimator of  $n!$ . Similarly,

$$\begin{aligned}
\mathbb{E}[z(n+1)^2] &= \sum_j P[c(n) = j] \mathbb{E}[z(n+1)^2 | c(n) = j] \\
&= P[c(n) = j] (2^{2(j+1)} (1/2^j) + 2^{2j} (1 - 1/2^j)) \\
&= P[c(n) = j] (2^{2j} - 2^j + 2^{j+2}) \\
&= P[c(n) = j] (2^{2j} + 3 \cdot 2^j) \\
&= \mathbb{E}[2^{2c(n)}] + 3\mathbb{E}[2^{c(n)}] \\
&= \mathbb{E}[z(n)^2] + 3n
\end{aligned} \tag{4}$$

Then, by iterating, we get

$$\begin{aligned}
\mathbb{E}[z(n)^2] &= 3(n-1) + \mathbb{E}[z(n-1)^2] \\
&= 3(n-1) + 3(n-2) + \mathbb{E}[z(n-2)^2] \\
&= \dots \\
&= 3(1 + 2 + \dots + (n-1)) + \mathbb{E}[z(0)^2] \\
&= 3n(n-1)/2 + 1
\end{aligned} \tag{5}$$

By Chebyshev's inequality, we can get the multiplicative error bound as follows:

$$\begin{aligned}
P[|z(n)/n| > \epsilon] &\leq \frac{\text{Var}[z(n)]}{\epsilon^2 \mathbb{E}[z(n)]^2} \\
&\leq \frac{\mathbb{E}[z(n)^2] - \mathbb{E}[z(n)]^2}{\epsilon^2 \mathbb{E}[z(n)]^2} \\
&= \frac{3n(n-1)/2 + 1 - n^2}{\epsilon^2 n^2} \\
&= O(1/\epsilon^2)
\end{aligned} \tag{6}$$

**Question.** Can we improve the error bound (sub-quadratic in  $\epsilon^{-1}$ )? How much will keeping around *multiple* counters of  $c$  help?

## 2 Counting Distinct Elements

Suppose we want to count  $d$  = number of *distinct* items from a stream of  $m$  elements where  $d \gg 1$ . The idea here is that we can use the maximum number of leading zeros ( $x$ ) to approximate the cardinality ( $\approx 2^x$ ).

### Algorithm

1. Let  $L = \lceil \log_2 m \rceil$ , and create an array of counters  $z[0 : L]$ .
2. Randomly pick  $h : [m] \rightarrow [N]$  from Pairwise Independent hash family, where  $N = 2^x - 1$  for some  $x \gg 1$ .
3. When an item  $a \in 1, \dots, m$  arrives, evaluate  $pos_a =$  number of trailing 0's in the binary representation of  $h(a)$  (i.e. largest  $j$  such that  $2^j | h(a)$ ).
4. Then,  $z[pos_a] += 1$
5. After the stream has passed, return  $\hat{d} = 2^k$ , where  $k$  is the largest number with  $z[k] > 0$ .

### Properties

1. Let  $X_{a,j} = 1$  if  $2^j$  is the largest power of 2 that divides  $h(a)$ , else  $X_{a,j} = 0$ . Then,  $P[X_{a,j} = 1] = 1/2^{j+1}$ , since  $h(a)$  would need to have  $1000 \dots 0$  ( $j$  0's) as the last  $j+1$  digits in its binary representation. Note that this is generally not true if  $N$  is small.
2. Let  $Y_j = \sum_{a: \text{distinct}} X_{a,j}$  (i.e. the number of distinct elements that increment counter  $j$ ). Then,  $\mathbb{E}[Y_j] = d/2^{j+1}$  by linearity of expectation. This means that  $Y_j \cdot 2^{j+1}$  is an unbiased estimator of  $d \forall j$ ! However, this does not mean that  $Y_j = z[j]$  in general. In fact, since  $z[j]$  can be incremented by duplicate items,  $Y_j \leq z[j] \forall j$ .
3.  $Y_j = 0$  iff  $z[j] = 0$ , since you can't have either of the values being greater than 0 without at least 1 (distinct) element incrementing them.
4.  $Var[Y_j] \leq \mathbb{E}[Y_j] \forall j$ , since the variation of [sum of pairwise independent indicator variables] (in this case,  $Y_j$ ) is always less than or equal to the expected value of the sum (this is true in general).

With these properties, we can bound the error for  $\hat{d}$ :

**Theorem.** With probability  $5/8$ ,  $d/16 \leq \hat{d} \leq 16d$ .

**Proof.** Let  $l$  be integer with  $2^l < d \leq 2^{l+1}$  and  $c$  some arbitrary constant. Then,

$$\begin{aligned}
P[\exists j \geq l + c : z[j] > 0] &\leq \sum_{j \geq l+c} P[z[j] > 0] (\because \text{union bound}) \\
&= \sum_{j \geq l+c} P[Y_j \geq 1] (\because \text{property 3}) \\
&\leq \sum_{j \geq l+c} \mathbb{E}[Y_j] (\because \text{Markov's Inequality}) \\
&= \sum_{j \geq l+c} d/2^{j+1} \\
&\leq \sum_{j \geq l+c} 2^{l+1}/2^{j+1} (\because d \leq 2^{l+1}) \\
&= 1/2^c \sum_{j \geq l+c} 1/2^{j-(l+c)} \\
&\leq 1/2^{c-1}
\end{aligned} \tag{7}$$

Thus,  $P[\mathbb{E}j \geq l + 4 : z[j] > 0] \leq 1/8$  (we simply plugged in  $c = 4$  to the above inequality). Hence, with probability at least  $7/8$ ,  $\hat{d} = 2^k \leq 2^{l+4} = 16d$ . This satisfies the right side of the theorem's inequality. Similarly,

$$\begin{aligned}
P[z[l - c] = 0] &= P[Y_{l-c} = 0] \\
&\leq P[|Y_{l-c} - \mathbb{E}[Y_{l-c}]| \geq \mathbb{E}[Y_{l-c}]] \\
&\leq \text{Var}[Y_{l-c}]/\mathbb{E}[Y_{l-c}]^2 (\because \text{Chebyshev}) \\
&\leq \mathbb{E}[Y_{l-c}]/\mathbb{E}[Y_{l-c}]^2 = 1/\mathbb{E}[Y_{l-c}] \\
&= 2^{l-c+1}/d \\
&< 2^{l-c+1}/2^l (\because d > 2^l \text{ from our definition of } l) \\
&= 1/2^{c-1}
\end{aligned} \tag{8}$$

Then, plugging in  $c = 3$  to the above inequality gives us  $P[z[l - 3] = 0] \leq 1/4$ . However,  $z[l - 3] = 0$  means  $k < l - 3$ , since  $k$  is supposed to be the biggest number that has nonempty slot in  $z$ . Thus, the negation of the inequality asserts that with probability at least  $3/4$ ,  $k \geq l - 3$ , i.e.  $\hat{d} = 2^k \geq 2^{l-3} = 2^{l+1}/2^4 \geq d/16$ . This satisfies the left side of the theorem's inequality. Thus, with probability  $1 - 1/4 - 1/8 = 5/8$  (these probabilities are the probability that either of the theorem's inequalities fail),  $d/16 \leq \hat{d} \leq 16d$ .  $\square$

**Question.** Can we tighten the bound on  $\hat{d}$  by keeping around *multiple* instances of  $z$ ? If not, how else could we improve the bound?